

Model Selection

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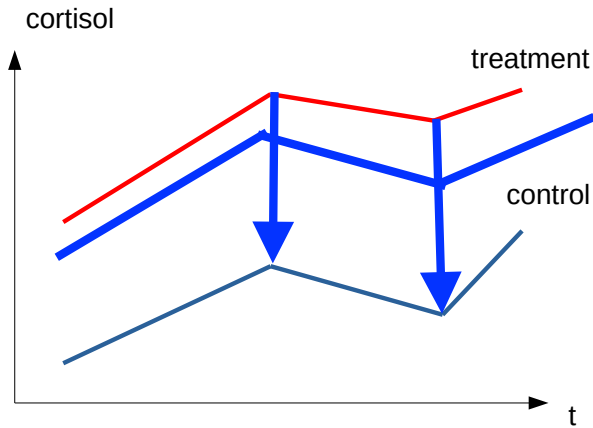
2 sides in breeding program:

1. economic evaluation

2. prediction of breeding values using statistical modelling

Why Statistical Modelling?

Some people believe, they do not need statistics. For them it is enough to look at a diagram



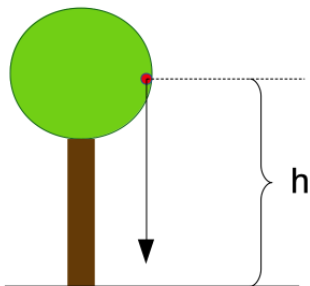
Statistical Modelling Because . . .

Two types of dependencies between physical quantities

1. deterministic  **no sources of uncertainty**
2. stochastic

Deterministic Versus Stochastic

deterministic



Law of gravity

time t that it takes the apple to fall can be computed without sources of uncertainty

stochastic

genotype

+

environment

many sources of uncertainty

> genes important

> post-translational processes

> env

phenotype

Statistical Model

- ▶ stochastic systems contains many sources of uncertainty
- ▶ statistical models can handle uncertainty
- ▶ components of a statistical model
 - ▶ response variable y **observed phenotypes**
 - ▶ predictor variables x_1, x_2, \dots, x_k **fixed effects and random breeding values**
 - ▶ error term e
 - ▶ function $m(x)$

statistical model

How Does A Statistical Model Work?

- ▶ predictor variables x_1, x_2, \dots, x_k are transformed by function $m(x)$ to explain the response variable y
- ▶ uncertainty is captured by error term.
- ▶ as a formula, for observation i

$$y_i = m(x_i) + e_i$$



x_i predictors are taken as input to the function $m()$

Which function $m(x)$?

- ▶ class of functions that can be used as $m(x)$ is infinitely large
- ▶ restrict to linear functions of predictor variables

Which predictor variables?


- ▶ Question, about which predictor variables to use is answered by model selection

predictor variables: x_1, x_2, \dots, x_k

Why Model Selection

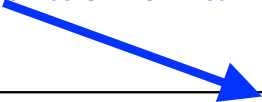
- ▶ Many predictor variables are available
- ▶ Are all of them relevant?
- ▶ What is the meaning of relevant in this context?

**with respect to explaining
the differences in the
response variables**



Example Dataset

random numbers = non-meaningful



Animal	Breast Circumference	Body Weight	RandPred
1	176	471	183
2	177	463	178
3	178	481	182
4	179	470	178
5	179	496	178
6	180	491	178
7	181	518	183
8	182	511	178
9	183	510	181
10	184	541	183

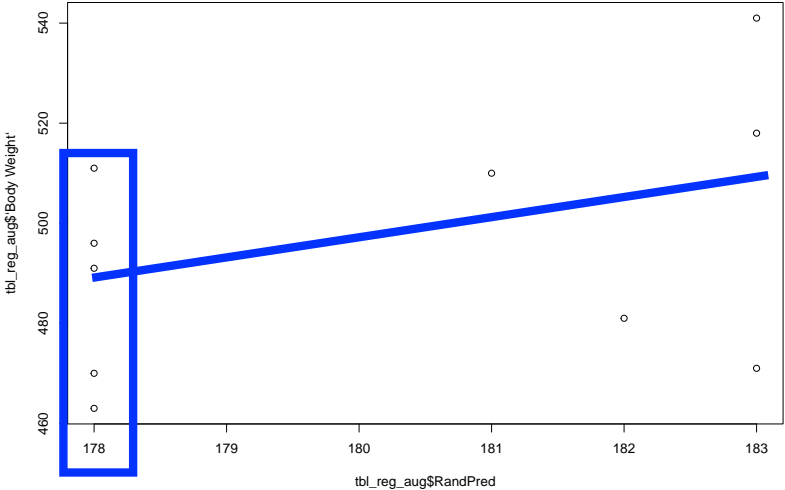
predictor = meaningful

response variable
 y_i



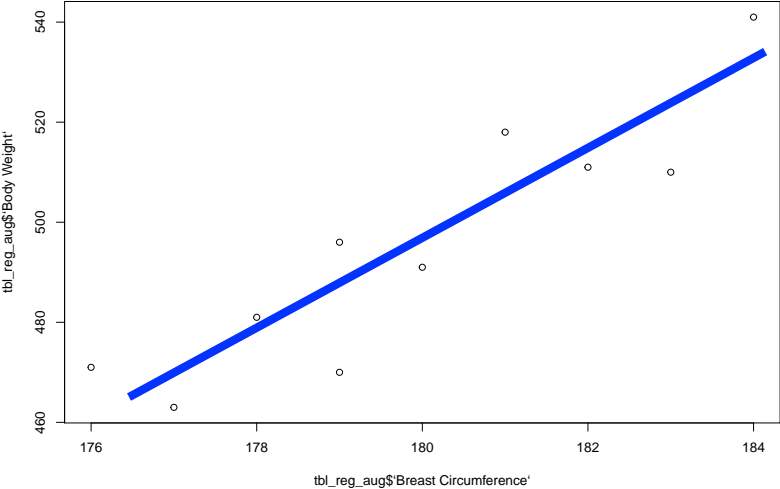
No Relevance of Predictors

random pattern



Relevance of Predictors

pattern: points all grouped around regression line



Fitting a Regression Model

```
##  
## Call:  
## lm(formula = `Body Weight` ~ RandPred, data = tbl_reg_aug)  
##  
## Residuals:  
##      Min       1Q   Median       3Q      Max   
## -35.574 -20.200   7.236  11.519  34.426   
##  
## Coefficients:  
##              Estimate Std. Error t value Pr(>|t|)      
## (Intercept) -236.775     608.880  -0.389   0.708      
## RandPred      4.062       3.379    1.202   0.264      
##  
## Residual standard error: 24.27 on 8 degrees of freedom  
## Multiple R-squared:  0.153, Adjusted R-squared:  0.04716   
## F-statistic: 1.445 on 1 and 8 DF,  p-value: 0.2636
```

**Low Adj. R-sq ==>
Model does not explain
a high rate of variation
of response variables**

std. error has about the same magnitude as the estimate ==> problem

Fitting a Regression Model II

```
##
## Call:
## lm(formula = `Body Weight` ~ `Breast Circumference`, data = tbl_reg_aug)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -17.3941  -6.5525  -0.0673   9.3707  13.2594
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)    -1065.115    255.483   -4.169 0.003126 **
## `Breast Circumference`     8.673     1.420    6.108 0.000287 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 11.08 on 8 degrees of freedom
## Multiple R-squared:  0.8234, Adjusted R-squared:  0.8014
## F-statistic: 37.31 on 1 and 8 DF, p-value: 0.000287
```

Multiple Regression

```
##  
## Call:  
## lm(formula = `Body Weight` ~ `Breast Circumference` + RandPred,  
##     data = tbl_reg_aug)  
##  
## Residuals:  
##      Min       1Q   Median       3Q      Max  
## -13.1363  -3.0404   0.7548   4.3149  14.3068  
##  
## Coefficients:  
##              Estimate Std. Error t value Pr(>|t|)  
## (Intercept)    -1492.865    295.360  -5.054 0.001473 **  
## `Breast Circumference`      8.304      1.202   6.909 0.000229 ***  
## RandPred        2.742        1.306   2.100 0.073839 .  
##  
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1  
##  
## Residual standard error: 9.278 on 7 degrees of freedom  
## Multiple R-squared:  0.8917, Adjusted R-squared:  0.8607  
## F-statistic: 28.81 on 2 and 7 DF,  p-value: 0.0004183
```

Which model is better?

Why not taking all predictors?

- ▶ Additional parameters must be estimated from data
- ▶ Predictive power decreased with too many predictors (cannot be shown for this data set, because too few data points)
- ▶ Bias-variance trade-off

Bias-variance trade-off

- ▶ Assume, we are looking for optimum prediction

Example with BW:
q_1 = {BC, RandPred}
q_2 = {BC}

$$s_i = \sum_{r=1}^q \hat{\beta}_{jr} x_{ijr}$$

with q relevant predictor variables

- ▶ Average mean squared error of prediction s_i

$$MSE = n^{-1} \sum_{i=1}^n E \left[(m(x_i) - s_i)^2 \right]$$

where $m(\cdot)$ denotes the linear function of the unknown true model.

Bias-variance trade-off II

the more predictors x are included in s_i ,
the smaller the bias will be

- ▶ MSE can be split into two parts

variance

$$MSE = n^{-1} \sum_{i=1}^n (E[s_i] - m(x_i))^2 + n^{-1} \sum_{i=1}^n \text{var}(s_i)$$

where $n^{-1} \sum_{i=1}^n (E[s_i] - m(x_i))^2$ is called the squared **bias**

- ▶ Increasing q leads to reduced bias but increased variance ($\text{var}(s_i)$)
- ▶ Hence, find s_i such that MSE is minimal
- ▶ Problem: cannot compute MSE because $m(\cdot)$ is not known

→ estimate MSE

Mallows C_p statistic

$$M: y_i = b_0 + b_1 * \text{breast_circumference}$$

► For a given model \mathcal{M} , $SSE(\mathcal{M})$ stands for the residual sum of squares.

► MSE can be estimated as

number of predictors in model M

$$\widehat{MSE} = n^{-1} SSE(\mathcal{M}) - \hat{\sigma}^2 + 2\hat{\sigma}^2 \frac{|\mathcal{M}|}{n}$$

where $\hat{\sigma}^2$ is the estimate of the error variance of the full model, $SSE(\mathcal{M})$ is the residual sum of squares of the model \mathcal{M} , n is the number of observations and $|\mathcal{M}|$ stands for the number of predictors in \mathcal{M}

$$C_p(\mathcal{M}) = \frac{SSE(\mathcal{M})}{\hat{\sigma}^2} - n + 2|\mathcal{M}|$$

Values of Mallows C_p should be as small as possible

Searching The Best Model

- ▶ Exhaustive search over all sub-models might be too expensive
- ▶ For p predictors there are $2^p - 1$ sub-models
- ▶ With $p = 16$, we get 6.5535×10^4 sub-models

→ step-wise approaches

2 ways to do step-wise approach:

- 1. forward selection**
- 2. backward elimination**

Forward Selection

M_0: $y_i = b_0 + e_i \implies$ just use an intercept $b_0 \implies$ compute C_p

Step 2: Question: would it be better to include any of the available pred?

\implies Constructing model M_1: In our example with BW \implies

Should M_1 contain BC or RandPred as its predictor?

1. Start with smallest sub-model \mathcal{M}_0 as current model
2. Include predictor that reduces SSE the most to current model
3. Repeat step 2 until all predictors are chosen

\rightarrow results in sequence $\mathcal{M}_0 \subseteq \mathcal{M}_1 \subseteq \mathcal{M}_2 \subseteq \dots$ of sub-models

4. Out of sequence of sub-models choose the one with minimal C_p

For each sub-model M_0, M_1, M_2, ... we have computed C_p

From all submodel select the one with lowest C_p value

This will be the best model.

Backward Selection

M_0: full model, for example of BW:

$$\mathbf{M}_0: y_i = b_0 + b_1 * \text{breast_circum.} + b_2 * \text{randpred} + e_i$$

Step 2: exclude predictors from M_0

1. Start with full model \mathcal{M}_0 as the current model
2. Exclude predictor variable that increases SSE the least from current model
3. Repeat step 2 until all predictors are excluded (except for intercept)

→ results in sequence $\mathcal{M}_0 \supseteq \mathcal{M}_1 \supseteq \mathcal{M}_2 \supseteq \dots$ of sub-models

4. Out of sequence choose the one with minimal C_p

Considerations

- ▶ Whenever possible, choose **backward** selection, because it leads to better results
- ▶ If $p \geq n$, only forward is possible, but then consider LASSO



backward elimination is not possible because full model cannot be fitted

Alternative Selection Criteria

AIC: Akaike Information Criterion

BIC: Bayesian Information Criterion

- ▶ AIC or BIC, requires distributional assumptions.
- ▶ AIC is implemented in `MASS::stepAIC()`
- ▶ Adjusted R^2 is a measure of goodness of fit, but sometimes is not conclusive when comparing two models
- ▶ Try in exercise

R-package: `olsrr`, `MASS` uses just AIC

For Genetic Evaluation:

- * **In our database: many different predictors are available for a given trait**
- * **Do model selection to find good balance between bias and variance**
- * **Model selection is used to identify fixed effects in our models to estimated variance components and to predict breeding values.**