## <span id="page-0-4"></span>**5.3 Pedigree BLUP**

In real datasets, the assumption of unrelated sires is unrealistic, because the selection process favors that male offspring of a given sire will be selected as sires again. As a consequence, the random sire efects are not independent. The dependence structure between the sire efects must be considered in the analysis. The sire model can still be written as shown in equation ([5.12\)](#page-0-0), but the variancecovariance matrix of the random sire efects (**s**) is no longer an identity matrix **I** times a common variance component  $\sigma_s^2$ , but it can be written as

$$
var(\mathbf{s}) = \mathbf{A}_s * \sigma_s^2 \tag{5.13}
$$

where  $\mathbf{A}_s$  is the sire relationship matrix. In a dataset with  $q$  sires, the matrix  $\mathbf{A}_s$  has dimensions  $q \times q$  and it contains the proportions of sire effects that are passed from father to son. Together with the sire variance  $\sigma_s^2$  this defines the variance-covariance structure of all sire efects. As an example, we can express the covariance  $cov(s_i, s_k)$  of the sire effects between son *i* and its sire k, as

$$
cov(s_i, s_k) = 1/2 * \sigma_s^2 \tag{5.14}
$$

where the factor  $1/2$  stems from the fact that sire k passes half of its genetic potential to its son  $i$ . Relating this single covariance back to the variancecovariance matrix  $A_s$  means that elements  $(i, k)$  and  $(k, i)$  are both  $1/2$ .

## <span id="page-0-2"></span>**5.3.1 Example Dataset**

<span id="page-0-1"></span>An application of the sire model is shown in the dataset given in Table [5.6](#page-0-1) which is taken from [\[Mrode](#page-0-2), [2005](#page-0-2)]

Table 5.6: Pre-weaning Gain in kg for fve beef animals

Animal	Sire	<b>Sex</b>	WWG
		M	4.5
5	3	F	2.9
6	1	F	3.9
		М	3.5
8	З	М	5.0

<span id="page-0-6"></span><span id="page-0-5"></span><span id="page-0-3"></span>The objective is to predict breeding values for sires 1, 3 and 4 based on the above dataset. The trait pre-weaning gain (WWG) is taken as response and Sex is assumed to be the only fxed efect. The following values for the variance components are assumed:  $\sigma_s^2 = 5$  and  $\sigma_e^2 = 55$ .

<span id="page-0-0"></span>This type of model where the structure of the variance-covariance matrix of the random effect is given by a pedigree cannot be fit by the package lme4. An

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extension of lme4 is given in the R-package pedigreemm. In pedigreemm it is possible to specify the variance-covariance structure via a pedigree. For our example of the dataset in Table [5.6](#page-0-1), this can be done as follows

```
library(pedigreemm)
ped_sire \leq pedigree(sire = c(rep(NA,2), 1), dam = rep(NA,3),
                   label = as-character(c(1,3,4))lmem_sire <- pedigreemm(
 formula = WWG ~ Sex + (1 | Sire),
 data = tbl_sire_model,
 pedigree = list(Sire = ped_sire)
)
(smry_lmem_sire <- summary(lmem_sire))
## Linear mixed model fit by REML ['lmerpedigreemm']
## Formula: WWG ~ Sex + (1 | Sire)
## Data: tbl_sire_model
##
## REML criterion at convergence: 8.5
##
## Scaled residuals:
## Min 1Q Median 3Q Max
## -1.1180 -0.6708 0.2236 0.6708 0.8944
##
## Random effects:
## Groups Name Variance Std.Dev.
## Sire (Intercept) 0.0000 0.0000
## Residual 0.5556 0.7454
## Number of obs: 5, groups: Sire, 3
##
## Fixed effects:
## Estimate Std. Error t value
## (Intercept) 3.4000 0.5270 6.451
## SexM 0.9333 0.6804 1.372
##
## Correlation of Fixed Effects:
## (Intr)
## SexM -0.775
```
The output of pedigreemm::pedigreemm() is equivalent to the one given by lme4::lmer(). From this we can see that the sire variance results in an estimate of 0. This does not correspond to the value that, we specifed as an assumption. This has two reasons. The assumed sire variance was not estimated from the small dataset in Table [5.6](#page-0-1), but from a larger dataset not shown here. The second reason is that it is not possible with pedigreemm::pedigreemm() to use an assumed variance component as input. The breeding values for the sires can be obtained by the function call

ranef(lmem\_sire)

## \$Sire ## (Intercept) ## 1 0 ## 3 0 ## 4 0

The predicted sire breeding values are also all equal to 0. The reason for this is that the sire variance component was estimated to be 0. Hence for small datasets, pedigreemm::pedigreemm() cannot be used.

## **5.3.2 Mixed Model Equations**

In a series of papers ([[Henderson](#page-0-3), [1953](#page-0-3)], [[Henderson](#page-0-4), [1963\]](#page-0-4) and [\[Henderson,](#page-0-5) [1975\]](#page-0-5)) which are summarized in [[Henderson](#page-0-6), [1982](#page-0-6)], a technique called **mixed model equations** was developed to solve for solutions of predicted values in a linear mixed effects model. For a given linear mixed effects model with  $var(e)$  =  $I * \sigma_e^2$  and  $var(s) = A_s * \sigma_s^2$ 

$$
y = Xb + Zs + e \tag{5.15}
$$

the solutions for the fxed efects estimates and the predicted values of the random efects can be obtained by solving the following set of equations.

$$
\begin{bmatrix}\nX^T X & X^T Z \\
Z^T X & Z^T Z + \lambda A_s^{-1}\n\end{bmatrix}\n\begin{bmatrix}\n\hat{b} \\
\hat{s}\n\end{bmatrix} =\n\begin{bmatrix}\nX^T y \\
Z^T y\n\end{bmatrix}
$$
\n(5.16)

where  $\lambda = \sigma_e^2/\sigma_s^2$ . For our example the matrices X and Z are defined as

$$
X = \begin{bmatrix} 0 & 1 \\ 1 & 0 \\ 1 & 0 \\ 0 & 1 \\ 0 & 1 \end{bmatrix}
$$

and

$$
Z = \begin{bmatrix} 1 & 0 & 0 \\ 0 & 1 & 0 \\ 1 & 0 & 0 \\ 0 & 0 & 1 \\ 0 & 1 & 0 \end{bmatrix}
$$

The matrix  $A_s^{-1}$  can be obtained by calling the function pedigreemm::getAinv() on the defned sire pedigree.

mat\_Ainv <- pedigreemm::getAInv(ped = ped\_sire)

That results in



The vector  $y$  corresponds to the vector of observations. With that we can solve the mixed model equations.

For the fxed efects we get

Table 5.7: Solutions for fxed Efect of Sex



For the random sire breeding values, we get

Table 5.8: Solutions for random Breeding Values of Sires

