

MLEM $\rightarrow p \gg N$ $\begin{cases} \rightarrow \text{MSE} \rightarrow \text{Shrinkage} \\ \rightarrow \text{same BLUE} \end{cases}$

LASSO $\rightarrow p \gg N \rightarrow$ reduction p , only important SNP

Least Absolute Shrinkage And Selection Operator (LASSO)

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Fixed Linear Effect Model

- ▶ Back to

$$y_i = \beta_0 + \sum_{j=1}^p \beta_j x_{ij} + \epsilon_i$$

- ▶ All $\beta_0, \beta_1, \dots, \beta_p$ into vector β of length $(p + 1)$

$$y = X\beta + \epsilon$$

$y = X\beta + \epsilon$ → back to FLEM

- ▶ Only random component: ϵ with

$$E(\epsilon) = 0 \text{ and } \text{var}(\epsilon) = I * \sigma^2$$

Parameter Estimation

- ▶ Least Squares

$$\hat{\beta}_{LS} = \underset{\beta}{\operatorname{argmin}} \|y - X\beta\|^2$$

(Handwritten red annotation: A bracket above the norm symbol $\|\cdot\|^2$ points to $\sum r_i^2$)

- ▶ Normal Equations

$$(X^T X) \hat{\beta}_{LS} = X^T y$$

(Handwritten red annotation: A bracket under $(X^T X)$ points to the text "solutions based on properties X")

- ▶ Existence of $(X^T X)^{-1}$?

1. Yes: $\hat{\beta}_{LS} = (X^T X)^{-1} X^T y$
2. No: $b_0 = (X^T X)^- X^T y$

with $(X^T X)^-$ being a generalized inverse of $(X^T X)$

Generalized Inverse

- ▶ System of equations

$$Ax = y$$

with coefficient matrix A , vector of unknowns x and vector of right hand side y

- ▶ If A^{-1} exists, then unknowns $x = A^{-1}y$
- ▶ If A^{-1} does not exist, $x = A^{-}y$ is one solution with A^{-} being a generalized inverse
- ▶ Generalized inverse A^{-} defined by

$$AA^{-}A = A$$

Solutions

- ▶ Why is A^{-} a solution
 - ▶ if $AA^{-}A = A$, then $AA^{-}Ax = Ax$
 - ▶ when $Ax = y$, this gives $A(A^{-}y) = y$
 - ▶ hence $A^{-}y = x$ is a solution
- ▶ If A^{-} is a generalized inverse of A then $Ax = y$ has solutions

$$\tilde{x} = A^{-}y + (A^{-}A - I)z$$

for arbitrary z

- ▶ Proof

$$A\tilde{x} = AA^{-}y + A(A^{-}A - I)z = AA^{-}y + (AA^{-}A - AI)z = AA^{-}y = y$$

because $AA^{-}A = A$.

Results

- ▶ $b_0 = (X^T X)^- X^T y$ is a solution to $(X^T X)b_0 = X^T y$
- ▶ But b_0 is not unique, because for any $(X^T X)^-$

$$\tilde{b}_0 = (X^T X)^- X^T y + ((X^T X)^- (X^T X) - I)z$$

is also a solution

- ▶ b_0 cannot be an estimate for β

Estimable Functions

Idea: construct linear functions ($q^T \beta$) of the parameters β such that

- ▶ estimator can be found from b_0
- ▶ independent of choice of b_0

Such linear functions $q^T \beta$ must satisfy

$$q^T \beta = t^T E(y)$$

for any vector t , then $q^T \beta$ is **estimable**

- ▶ Determine q as

$$q^T = t^T X$$

Invariance to b_0

When $q^T \beta$ is estimable, then

- ▶ $q^T b_0$ is always the same, independent of choice of b_0
- ▶ Why?
- ▶ With $q^T = t^T X$

$$q^T b_0 = t^T X b_0 = t^T X (X^T X)^- X^T y$$

is independent of choice of b_0 because $X(X^T X)^- X^T$ is independent of choice of $(X^T X)^-$

Summary

Use of generalized inverse $(X^T X)^-$ of normal equations yields

- ▶ solutions b_0
- ▶ estimable functions $q^T b_0$ which estimate $q^T \beta$
- ▶ independent of b_0

But for genomic data

- ▶ no possibility to determine important SNP loci
- ▶ need an alternative to least squares

Alternatives To Least Squares

Desirable properties

1. **Subset Selection:** determine important predictors
2. **Shrinkage:** limit parameter estimates to certain area
3. **Dimension Reduction:** Reduce p predictors to m linear combinations where $m < p$

→ reduce p

~~X~~

LASSO

- ▶ ... stands for Least Absolute Shrinkage and Selection Operator
- ▶ ... combines subset selection (1) and shrinkage (2)
- ▶ shrinkage is achieved by introduction of penalty term
- ▶ subset selection is due to the form of penalty term

Shrinkage

- ▶ penalty term added to least squares criterion

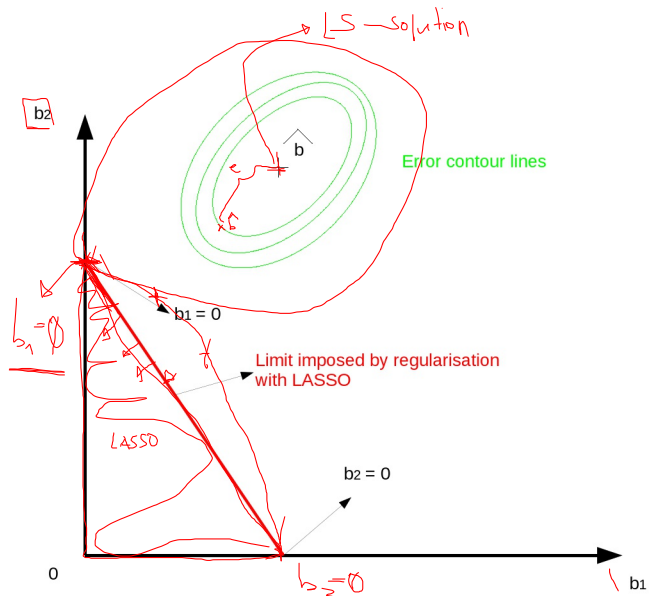
$$\hat{\beta}_{LASSO} = \underset{\beta}{\operatorname{argmin}} \left\{ \sum_{i=1}^n \left(y_i - \beta_0 - \sum_{j=1}^p \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^p |\beta_j| \right\}$$

penalty & min

$\beta_j = 0$

- ▶ large values of $|\beta_j|$ are penalized compared to small $|\beta_j|$

Subset Selection



Find λ

- weight for penalties*
- ▶ λ is an additional parameter to be estimated from data
 - ▶ use cross validation *CV*
 - ▶ split data randomly into training set (80 – 90%) and test set (10 – 20%)
 - ▶ assume a certain λ value and do parameter estimation with training data *$\hookrightarrow \lambda=1$*
 - ▶ try to predict test data with estimated parameters *\Rightarrow residual*
 - ▶ repeat this many times
 - ▶ take that λ with the best predictive performance