Variance Components Estimation

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Genetic Variation

- \triangleright Requirement for trait to be considered in breeding goal
- \triangleright Breeding means improvement of next generation via selection and mating
- ▶ Only genetic (additive) components are passed to offspring
- \triangleright Selection should be based on genetic component of trait
- \triangleright Selection only possible with genetic variation

 \rightarrow genetic variation indicates how good characteristics are passed from parents to offspring

 \rightarrow measured by **heritability** $h^2 = \frac{\sigma_{\theta}^2}{\sigma_{\rho}^2}$

Two Traits

Problems

- ▶ Genetic components cannot be observed or measured
- \blacktriangleright Must be estimated from data
- \blacktriangleright Data are mostly phenotypic
- \rightarrow topic of variance components estimation
	- \blacktriangleright Model based, that means connection between phenotypic measure and genetic component are based on certain model

$$
p=g+e
$$

with $cov(g, e) = 0$

Goal: separate variation due to $g(\sigma_a^2)$ from phenotypic variation

Example of Variance Components Separation

- \blacktriangleright Estimation of repeatability
- \triangleright Given repeated measurements of same trait at the same animal
- \blacktriangleright Repeatability means variation of measurements at the same animal is smaller than variation between measurements at different animals

Repeatability Plot

Model

$$
y_{ij} = \mu + t_i + \epsilon_{ij}
$$

where

- y_{ij} measurement *j* of animal *i*
- μ expected value of y
- t_i deviation of y_{ij} from μ attributed to animal i
- ϵ_{ij} measurement error

Estimation Of Variance Components

$$
\blacktriangleright E(t_i)=0
$$

 $\sigma_t^2 = E(t_i^2)$: variance component of total variance (σ_y^2) which can be attributed to the t-effects

$$
\blacktriangleright E(\epsilon_{ij})=0
$$

 $\blacktriangleright \sigma_{\epsilon}^2 = E(\epsilon_{ij}^2)$: variance component attributed to ϵ -effects

$$
\blacktriangleright \; \sigma_y^2 = \sigma_t^2 + \sigma_{\epsilon}^2
$$

 \blacktriangleright Repeatability w defined as:

$$
w = \frac{\sigma_t^2}{\sigma_t^2 + \sigma_\epsilon^2}
$$

 \rightarrow estimate of σ_t^2 needed

Analysis Of Variance (ANOVA)

where

$$
SSQ(t) = \left[\frac{1}{n}\sum_{i=1}^{r} \left(\sum_{j=1}^{n} y_{ij}\right)^{2}\right] - \left(\sum_{i=1}^{r} \sum_{j=1}^{n} y_{ij}\right)^{2} / N
$$

$$
SSQ(\epsilon) = \sum_{i=1}^{r} \sum_{j=1}^{n} y_{ij}^{2} - \left[\frac{1}{n} \sum_{i=1}^{r} \left(\sum_{j=1}^{n} y_{ij}\right)^{2}\right]
$$

Zahlenbeispiel

Df Sum Sq Mean Sq F value Pr(>F) ## Bull 9 286.7 31.85 13.85 8.74e-07 *** ## Residuals 20 46.0 2.30 ## --- ## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Setting expected values of Mean Sq equal to estimates of variance components

$$
\hat{\sigma}_{\epsilon}^2 = 2.3
$$
 and $\hat{\sigma}_{t}^2 = \frac{31.85 - 2.3}{3} = 9.85$

Repeatability

$$
\hat{w} = \frac{\hat{\sigma}_t^2}{\hat{\sigma}_t^2 + \hat{\sigma}_\epsilon^2} = 0.81
$$

Same Strategy for Sire Model

 \triangleright Sire model is a mixed linear effects model with sire effects s as random components

$$
y = Xb + Zs + e
$$

In case where sires are not related, $\text{Svar(s)} = I * \sigma_s^2$ From σ_s^2 , we get genetic additive variance as $\sigma_a^2 = 4 * \sigma_s^2$

ANOVA

with

$$
k = \frac{1}{r-1} \left[N - \frac{\sum_{i=1}^{r} n_i^2}{N} \right]
$$

Maximum Likelihood (ML)

\blacktriangleright Likelihood

$$
L(\theta) = f(y|\theta)
$$

\blacktriangleright Normal distribution

$$
L(\theta) = (2\pi)^{-1/2n} \sigma^{-n} |H|^{-1/2} * \exp \left\{-\frac{1}{2\sigma^2} (y - Xb)^T H^{-1} (y - Xb)\right\}
$$

with $var(y) = H * \sigma^2$ and $\theta^T = \begin{bmatrix} b & \sigma^2 \end{bmatrix}$

Maximization of Likelihood

$$
\blacktriangleright \ \mathsf{Set} \ \lambda = \mathsf{log} L
$$

I Compute partial derivatives of λ with respect to all unknowns

$$
\frac{\partial \lambda}{\partial b}
$$

$$
\frac{\partial \lambda}{\partial \sigma^2}
$$

 \triangleright Set partial derivatives to 0 and solve for unknowns \blacktriangleright Use solutions as estimates

Restricted Maximum Likelihood (REML)

► Problem with ML: estimate of σ^2 depends on $b \to$ undesirable

 \triangleright Do transformations Sy and Qy

- (i) The matrix S has rank $n t$ and the matrix Q has rank t
- (ii) The result of the two transformations are independent, that means $cov(Sy, Qy) = 0$ which is met when $SHQ^T = 0$
- (iii) The matrix S is chosen such that $E(Sy) = 0$ which means $SX = 0$
- (iv) The matrix QX is of rank t, so that every linear function of the elements of Qy estimate a linear function of b .

From (i) and (ii) it follows that the likelihood L of y is the product of the likelihoods of Sy (L^*) and Qy (L^{**}) that means

$$
\lambda=\lambda^*+\lambda^{**}
$$

► Variance components are estimated from λ^* which will then be independent of b

Bayesian Estimation

- \blacktriangleright Proposed already in the 80's
- \blacktriangleright Full implementation only in 1993
- \blacktriangleright Requirements:
	- \blacktriangleright cheap computing and
	- \triangleright good pseudo-random number generators
- \triangleright Bayesian estimation is based on conditional posterior distribution of unknowns given the knowns
- \triangleright Conditional posterior distribution is computed from prior distribution of unknowns times the likelihood

Model

 \blacktriangleright Univariate Gaussian linear mixed model

 $y = Xb + Zu + e$

where

- y vector of observations (length *n*)
- b vector of fixed effects (length p)
- u vector of random breeding values (length q)
- e vector of random residuals (length n)
- X $n \times p$ design matrix linking fixed effects to observations
- Z $n \times q$ design matrix linking breeding values to observations

 \blacktriangleright Data generating distribution

$$
y|b, u, \sigma_e^2 \sim \mathcal{N}(Xb + Zu, I * \sigma_e^2)
$$

where I is a $n \times n$ identity matrix and $\sigma_{\sf e}^2$ is the variance of the random residuals.

Priors

 \blacktriangleright Prior distributions must be specified for all unknowns **I** Unknowns in our example are: *b*, *u*, σ_e^2 and σ_u^2 \blacktriangleright Prior distribution for \triangleright b is flat, i.e. $p(b) \propto c$ ► *u* Normal distribution as $u|G, \sigma_u^2 \sim N(0, G * \sigma_u^2)$ \triangleright σ_e^2 scaled inverse χ^2 : $p(\sigma_e^2|\nu_e,\mathsf{s}_e^2) \propto (\sigma_e^2)^{-\nu_e/2-1} \text{exp}(-\frac{1}{2}\nu_e\mathsf{s}_e^2/\sigma_e^2)$ $\sigma_u^2 : p(\sigma_u^2 | \nu_u, s_u^2) \propto (\sigma_u^2)^{-\nu_u/2 - 1} \exp(-\frac{1}{2}\nu_u s_u^2/\sigma_u^2)$ ν_e , ν_s , s_e^2 and s_u^2 are called hyper-parameters and must be determined

Additional Terms

$$
\theta^{\mathcal{T}} = (b^{\mathcal{T}}, u^{\mathcal{T}}) = (\theta_1, \theta_2, \dots, \theta_N)
$$

$$
\theta_{-i} = (\theta_1, \theta_2, \dots, \theta_{i-1}, \theta_{i+1}, \dots, \theta_N)
$$

$$
\blacktriangleright
$$
 Further, let

$$
s^\mathcal{T}=(s_u^2,s_e^2)
$$

and

$$
\nu^{\mathcal{T}} = (\nu_u, \nu_e)
$$

Joint Posterior Density

The joint posterior distribution can be written as

$$
p(\theta, \sigma_u^2, \sigma_e^2 | y, s, \nu) \propto p(\theta) * p(\sigma_u^2 | \nu_u, s_u^2) * p(\sigma_e^2 | \nu_e, s_e^2) * p(y | \theta, \sigma_e^2)
$$

Fully Conditional Posterior Densities of *θ*

 \triangleright Density of every single unknown component when setting all other components as known

$$
\theta_i|y, \theta_{-i}, \sigma_u^2, \sigma_e^2, s, \nu \sim \mathcal{N}(\tilde{\theta}_i, \tilde{v}_i)
$$

where
$$
\tilde{\theta}_i = (r_i - \sum_{j=1, j \neq i}^{N} w_{ij} \theta_j) / w_{ii}
$$
 and $\tilde{v}_i = \sigma_e^2 / w_{ii}$.

riangleright vector r is the vector of right-hand side of MME \blacktriangleright matrix W is the coefficient matrix of MME

Fully Conditional Posterior Densities of *σ* 2 e

I scaled inverted chi-square distribution for σ_e^2

$$
\sigma_{\rm e}^2|y,\theta,\sigma_u^2,s,\nu\sim\tilde{\nu_{\rm e}}\tilde{s_{\rm e}}^2\chi_{\tilde{\nu_{\rm e}}}^{-2}
$$

 \blacktriangleright Parameters of the above distribution are defined as

$$
\tilde{\nu_e} = n + \nu_e
$$

and

$$
\tilde{s}_{e}^{2} = \left[\left(y - Xb - Zu \right)^{T} \left(y - Xb - Zu \right) + \nu_{e} s_{e}^{2} \right] / \tilde{\nu}_{e}
$$

Fully Conditional Posterior Densities of *σ* 2 u

 \blacktriangleright scaled inverted chi-square distribution for σ_u^2

$$
\sigma_u^2|y,\theta,\sigma_e^2,s,\nu\sim\tilde{\nu_u}\tilde{s_u}^2\chi_{\tilde{\nu_u}}^{-2}
$$

 \blacktriangleright Parameters of the above distribution are defined as

$$
\tilde{\nu_u}=q+\nu_u
$$

and

$$
\tilde{s_u}^2 = \left[u^T G^{-1} u + \nu_u s_u^2 \right] / \tilde{\nu_u}
$$

Implementation

- Step 1: set starting values for θ , σ_e^2 and σ_u^2
- **If** Step 2: draw random number for each component θ_i of θ from fully conditional distribution $\mathcal{N}(\widetilde{\theta}_i, \widetilde{\mathsf{v}}_i)$
- Step 3: draw random number for σ_e^2 from $\tilde{\nu}_e \tilde{s}_e^2 \chi_{\tilde{\nu}_e}^{-2}$
- ► Step 4: draw random number for σ_u^2 from $\tilde{\nu}_u \tilde{s}_u^2 \chi_{\tilde{\nu}_u}^{-2}$
- Repeat steps 2-4 many times and store random numbers
- \triangleright Step 5: compute means of random numbers to get Bayesian estimates of unknowns θ , $\sigma_{\bm{e}}^2$ and $\sigma_{\bm{u}}^2$